

Connections between Calibration and Statistics with illustrations from RGS-pn rectification

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An XMM-Newton RGS spectrum





3C273 with the 2009 XMM calibrations & SAS v9





RGS-pn rectification



Action 2009-05-07/03: The Instrument Teams should establish a time epoch-dependent fudge function for the RGS effective area such that the joint analysis of RGS and EPIC data is possible. Results of this effort should be presented at the next UG meeting in 2010 with the aim to make it available to the general user after the review.

This action item shall not prevent the instrument teams to continue their studies of the crosscalibration. It shall be a SAS 'working package' for the users allowing them simultaneous fits when needed. In any case, remaining uncertainties in the calibration will need to be well and clearly documented.

- XCal sample of ~50 RGS≒EPIC spectra and models
 - 3C273, PKS2155-304, H1426+428, PKS0548-322, Mkn501, Mkn180, 1H1219+301
- Adopt XCal methods
 - Pile-up
 - χ²
 - 25 or more counts per bin
 - XCal model parameter constraints
- Rectify XCal RGS models with XSPEC user model rgsrectify
 - 33 rectification factors {R6, R7, R8,...., R36, R37, R38} in $\Delta\lambda$ =±0.5Å
 - RGS1 & RGS2

RGS-pn rectification at the 2010 Users Group



European Space Agency

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Statistical benefits of RGS-pn rectification





Statistical benefits of RGS-pn oxyfication





RGS-pn rectification in practice



- Contributions to RGS-pn rectification factors
 - calibration systematics
 - RGS
 - effective area including instrumental oxygen
 - EPIC-pn
 - effective area
 - redistribution
 - PSF
 - data analysis systematics
 - physical model inaccuracies including interstellar oxygen
 - pile-up
- How to rectify in SAS v10 by RGS RMF modification
 - rgsproc ... withrectification=yes
 - not an RGS effective area correction

SRN269 final CCF rectification result



	$7.5 < \lambda(\text{\AA}) \le 23.5$	$23.5 < \lambda(Å) \le 37.5$
RGS1	0.9716	1.0021
RGS2	0.9753	1.0028

Analysis in high-energy astrophysics



data ⇔ models

 $\{n_i\}_{i=1,N} \Leftrightarrow \{\mu_i\}_{i=1,N}$

 \ge 0 individual events \Leftrightarrow continuously distributed

detector coordinates \Leftrightarrow physical parameters

never change \Leftrightarrow change limited only by physics

have no errors \Leftrightarrow subject to fluctuations

most precious resource \Leftrightarrow predictions possible

kept forever in archives \Leftrightarrow kept forever in journals and textbooks



Likelihood of data on models



$$\{n_{i}\}_{i=I,N} \text{ data} \text{ statistics models } \{\mu_{i}\}_{i=I,N}$$

$$L = \prod_{i=1}^{N} P(n_{i} \mid \mu_{i})$$
Gaussian

$$L = \prod_{i=1}^{N} \frac{1}{\sigma_{i}\sqrt{2\pi}} \exp\left(-\frac{(n_{i} - \mu_{i})^{2}}{2\sigma_{i}^{2}}\right) dn_{i}$$

$$L = \prod_{i=1}^{N} \frac{e^{-\mu_{i}}\mu_{i}^{n_{i}}}{n_{i}!}$$

$$\ln L = -\frac{1}{2}\sum_{i=1}^{N} \frac{(n_{i} - \mu_{i})^{2}}{\sigma_{i}^{2}} - \sum_{i=1}^{N} \ln \sigma_{i} + \kappa(\ln dn_{i})$$

$$\ln L = \sum_{i=1}^{N} n_{i} \ln \mu_{i} - \mu_{i} - \kappa(\ln n_{i}!)$$

$$-2\ln L = \chi^{2}$$

$$-2\ln L = C$$
Cash 1979, ApJ, 228, 939

Trivial maximum-likelihood solution



What is the average number of counts per bin ?

$$\ln L = \sum_{i=1}^{N} n_{i} \ln \mu - \mu \qquad \ln L = \sum_{i=1}^{N} \frac{(n_{i} - \mu)^{2}}{n_{i}} \qquad \ln L = \sum_{i=1}^{N} \frac{(n_{i} - \mu)^{2}}{\mu}$$
$$\frac{\partial \ln L}{\partial \mu} = \sum_{i=1}^{N} \frac{n_{i}}{\mu} - 1 = 0$$
$$\sum_{i=1}^{N} n_{i} = \mu \sum_{i=1}^{N} 1$$

Bias



Maximum-likelihood estimates, μ , of the mean counts for observations $\{n\}$

- > χ^2 data weights $\mu^{-1} = \langle n^{-1} \rangle$
- > C-statistic $\mu = \langle n \rangle$ (the correct answer)
- > χ^2 model weights $\mu^2 = \langle n^2 \rangle$

Biases for Poisson distribution with $\mu = 100$

- > $1/\langle n^{-1} \rangle = 98.9897$
- > $\sqrt{<n^2>}$ = 100.4988
- Bias is binning dependent
- Unbias is binning independent

RGS-pn rectification alternatives



XSPEC statistic	RGS1	RGS2	RGS1	RGS2
χ^2 (data)	-2.8%	-2.7%	+0.1%	+0.2%
С	-0.4%	-0.2%	+3.9%	+3.3%
χ^2 (model)	+1.2%	+1.5%	+5.0%	+5.6%
λ	short	short	long	long

Here the choice of statistical method makes a difference.

Gaussian or Poisson ?



- The choice
 - XSPEC> statistic chisq
 - XSPEC> statistic cstat
- For high counts they are nearly the same $(\sigma^2 = n)$
- Gaussian chisq
 - the default
 - the wrong answer
 - asymptotic goodness-of-fit
 - rebin to "improve the statistics" or "avoid low-count bias"
 - $n \ge 5$ or 10 or 25 or 100 according to taste
- Poisson cstat
 - the correct answer for all $n \ge 0$
 - no rebinning necessary
 - asymptotic goodness-of-fit

To rebin or not to rebin a spectrum ?



Pros

- Gaussian \equiv Poisson for $n \gg 0$
- dangers of oversampling
- saves time
- everybody does it
- "improves the statistics"
- grppha and other tools exist
- on log-log plots $ln0=-\infty$

• Cons

- rebinning throws away information
- 0 is a perfectly good measurement
- images are never rebinned
- Poisson statistics robust for $n \ge 0$
- $\mu_1 + \mu_2$ is also a Poisson variable
- oversampling harmless

Leave spectra alone. Don't rebin. Use Poisson statistics.

10 commandments of data analysis



- > Use all the data at your disposal
- Don't alter data
- Make the model as complete as possible
- Use the most accurate statistics
- Support decisions with unreduced statistics
- Report parameter estimates and errors
- Beware of upper limits
- Be aware of systematic errors
- Make informative unbiased plots
- > Distinguish physical and instrumental coordinates

10 commandments of IACHEC data analysis



- Don't rebin spectra
- n=0 is a perfectly good measurement
- Don't subtract from the data, add to the model
- Use the C-statistic
- Report unreduced C-statistic, NBINS & NDOF (and NFREE/NPAR)
- > Report maximum-likelihood parameter estimates and $\Delta C=1$ errors
- > $\mu = 0.\pm \sigma$ is a perfectly good estimate
- Beware of systematic errors
- Beware of log-log plots
- Beware of PI redistribution